

Topic 12

Overview of Estimation

Classical Statistics

Outline

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Classical Statistics

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Introduction

In the simplest possible terms, the goal of **estimation theory** is to answer the question:

What is that number?

Statistics has provided two distinct approaches this question - typically called

- **classical** or frequentist, and
- **Bayesian**.

Definition. A **statistic** is a function of the data that does not depend on any unknown parameter.

Exercise. Give a listing of statistics seen to this point.

Parameter Estimation

For **parameter estimation**, we consider $X = (X_1, \dots, X_n)$, independent random variables chosen according to one of a family of probabilities P_θ where θ is element from the **parameter space** Θ . Based on our analysis, we choose an **estimator** $\hat{\theta}(X)$. If the **data** \mathbf{x} takes on the values x_1, x_2, \dots, x_n , then

$$\hat{\theta}(x_1, x_2, \dots, x_n)$$

is called the **estimate** of θ . Thus we have three closely related objects.

1. θ - the **parameter**, an element of the parameter space, is a number or a vector.
2. $\hat{\theta}(x_1, x_2, \dots, x_n)$ - the **estimate**, is a number or a vector obtained by evaluating the estimator on the data $\mathbf{x} = (x_1, x_2, \dots, x_n)$.
3. $\hat{\theta}(X_1, \dots, X_n)$ - the **estimator**, is a random variable. We will analyze the distribution of this random variable to decide how well it performs in estimating θ .

Parameter Estimation

For **Bernoulli trials** $X = (X_1, \dots, X_n)$, we have

1. p , a single parameter, the **probability of success**, with parameter space $[0, 1]$.
2. $\hat{p}(x_1, \dots, x_n)$ is the **sample proportion** of successes in the data set.
3. $\hat{p}(X_1, \dots, X_n)$, the **sample mean** of the random variables

$$\hat{p}(X_1, \dots, X_n) = \frac{1}{n}(X_1 + \dots + X_n) = \frac{1}{n}S_n$$

is an estimator of p . We can give the distribution of this estimator because S_n is a **binomial** random variable.

Parameter Estimation

Given pairs of observations $(\mathbf{x}, \mathbf{y}) = ((x_1, y_1), (x_2, y_2), \dots, (x_n, y_n))$ that display a general linear pattern, we use **ordinary least squares regression** for

1. parameters - the **slope** β and **intercept** α of the regression line. So, the parameter space is \mathbb{R}^2 , pairs of real numbers.
2. They are estimated using the **statistics** $\hat{\beta}$ and $\hat{\alpha}$ in the equations

$$\hat{\beta}(\mathbf{x}, \mathbf{y}) = \frac{\text{cov}(\mathbf{x}, \mathbf{y})}{\text{var}(\mathbf{x})}, \quad \bar{y} = \hat{\alpha}(\mathbf{x}, \mathbf{y}) + \hat{\beta}(\mathbf{x}, \mathbf{y})\bar{x}.$$

3. Later, when we consider statistical inference for linear regression, we will analyze the distribution of the estimators.

Classical Statistics

In classical statistics, the **state of nature** is assumed to be fixed, but unknown to us. Thus, one goal of estimation is to determine which of the P_θ is the source of the data. The **estimate** is a statistic

$$\hat{\theta} : \text{data} \rightarrow \Theta.$$

For estimation procedures, the classical approach to statistics is based on two fundamental questions:

- How do we determine estimators?
- How do we evaluate estimators?
 - Does this estimator in any way systematically under or over estimate the parameter?
 - Does it has large or small variance?
 - How does it compare to a notion of best possible estimator?
 - How easy is it to determine and to compute?
 - How does the procedure improve with increased sample size?

Densities and Likelihoods

Our analysis is based on the **distribution of the random variables** that underlie the data under any value θ . For each $\theta \in \Theta$, we have a **density function**

$$f_X(\mathbf{x}|\theta).$$

For experimental designs based on a **simple random sample**, the observations X_1, \dots, X_n , are drawn from a family of distributions each having density $f_X(x|\theta)$. For **independent** random variables, the **joint density** is the **product** of the **marginal densities**

$$f_X(\mathbf{x}|\theta) = \prod_{k=1}^n f_X(x_k|\theta) = f_X(x_1|\theta)f_X(x_2|\theta) \cdots f_X(x_n|\theta).$$

In this circumstance, the data \mathbf{x} are **known** and the parameter θ is **unknown**. Thus, we write the density function as

$$L(\theta|\mathbf{x}) = f_X(\mathbf{x}|\theta)$$

and call L the **likelihood function**.

Densities and Likelihoods

- For **Bernoulli trials** with a known number of trials n but unknown success probability parameter p has joint density

$$\begin{aligned} \mathbf{f}_X(\mathbf{x}|p) &= p^{x_1}(1-p)^{1-x_1} p^{x_2}(1-p)^{1-x_2} \dots p^{x_n}(1-p)^{1-x_n} = p^{\sum_{k=1}^n x_k} (1-p)^{\sum_{k=1}^n (1-x_k)} \\ &= p^{\sum_{k=1}^n x_k} (1-p)^{n-\sum_{k=1}^n x_k} = p^{n\bar{x}} (1-p)^{n(1-\bar{x})} \end{aligned}$$

- Normal random variables** unknown mean μ and standard deviation σ has joint density

$$\begin{aligned} \mathbf{f}_X(\mathbf{x}|\mu, \sigma) &= \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x_1 - \mu)^2}{2\sigma^2}\right) \dots \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x_n - \mu)^2}{2\sigma^2}\right) \\ &= \frac{1}{(\sigma\sqrt{2\pi})^n} \exp\left(-\frac{1}{2\sigma^2} \sum_{k=1}^n (x_k - \mu)^2\right) \end{aligned}$$

Exercise. Find the joint density of n independent $\Gamma(\alpha, \beta)$ random variables.