

# Topic 19

## Extensions on the Likelihood Ratio

### Chi-square Tests

# Outline

Overview

Chromosomal Recombination

# Overview

For the two-sample two-sided sample proportion test, we used the **central limit theorem** to assert that our test statistic is approximately the square of a standard normal random variable, and hence is a  $\chi^2$  random variable with **1 degree of freedom**.

These ideas can be extended to the case in which  $\Theta$  is a  **$d$ -dimensional parameter space** and  $k$  of these parameters are, under the null hypothesis, assume to have fixed values. Thus,  $\Theta_0$  is  $d - k$ -dimensional.

As before

- $\hat{\theta}_0$  be the parameter value that **maximizes** the likelihood for  $\theta \in \Theta_0$  and
- $\hat{\theta}$  be the parameter value that **maximizes** the likelihood for  $\theta \in \Theta$ .

## Procedure

**Theorem.** Whenever the **maximum likelihood estimator** has an asymptotically normal distribution, let  $\Lambda(\mathbf{x})$  be the **likelihood ratio**

$$\Lambda(\mathbf{x}) = \frac{L(\hat{\theta}_0|\mathbf{x})}{L(\hat{\theta}|\mathbf{x})}.$$

for an  **$d$ -dimensional** parameter space. Consider the hypothesis:

$$H_0 : \theta_i = c_i \text{ for all } i = 1, \dots, k \quad \text{versus} \quad H_1 : \theta_i \neq c_i \text{ for some } i = 1, \dots, k.$$

Then under  $H_0$ , the distribution of

$$-2 \ln \Lambda_n(X)$$

converge to a  $\chi_k^2$  distribution as the sample size  $n \rightarrow \infty$ .

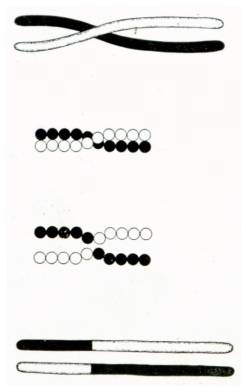
**NB.** More generally, if  $\Theta_0$  is defined has by  $k$  linear **constraints**, this can be changed to the test above by a linear change of variables. Thus,

$$\text{degrees of freedom} = \dim(\Theta) - \dim(\Theta_0).$$

## Chromosomal Recombination

During **meiosis**, paired chromosomes experience **crossing over** events in the formation of gametes. During **prophase I**, the four available chromatids (two from each parent) are aligned allowing breaks and reattachments of homologous sites.

Recombination can occur with a small probability at anywhere along chromosome. The number of crossing over events can be modeled as **Poisson random variable**. The mean number of cross overs for a given chromosomal segment is called its **genetic length** with **Morgans** as the unit of measurement.



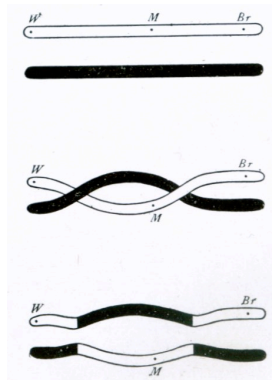
**Thomas Morgan's** 1916 drawing of describing a chromosomal crossing over event.

## Chromosomal Recombination

With whole **genome sequences for trios**, an individual along with both parents, we can determine the number of crossing over events on both parents and address the question: *Are the number of crossing over events different in sperm and in eggs?* This leads to the hypothesis

$$H_0 : \lambda_m = \lambda_f \quad \text{versus} \quad H_1 : \lambda_m \neq \lambda_f$$

where  $\lambda_m$  and  $\lambda_f$  are the **parameters** (male and female) in the **Poisson random variable** that gives the number of crossing over events in the human chromosome across all **22 autosomes**.



**Thomas Morgan's** 1916 drawing of a double crossing over event.

## Chromosomal Recombination

The **data** are  $n_m$  and  $n_f$  the number of crossing over events for each parent's (male's and female's) chromosome. Thus, assuming that the recombination sites are independent on the two parents, the **likelihood function** is

$$L(\lambda_m, \lambda_f | n_m, n_f) = \frac{\lambda_m^{n_m}}{n_m!} e^{-\lambda_m} \cdot \frac{\lambda_f^{n_f}}{n_f!} e^{-\lambda_f}.$$

**Exercise.** Show that the **maximum likelihood estimates** for the likelihood function above is

$$\hat{\lambda}_m = n_m \quad \text{and} \quad \hat{\lambda}_f = n_f.$$

Thus,

$$L(\hat{\lambda}_m, \hat{\lambda}_f | n_m, n_f) = \frac{n_m^{n_m}}{n_m!} \cdot \frac{n_f^{n_f}}{n_f!} e^{-(n_m+n_f)}.$$

## Chromosomal Recombination

Under the **null hypothesis**,  $\lambda_m$  and  $\lambda_f$  have a common value. Let's denote this by  $\lambda_0$ . Then the likelihood function is

$$L(\lambda_0 | n_m, n_f) = \frac{\lambda_0^{n_m}}{n_m!} e^{-\lambda_0} \cdot \frac{\lambda_0^{n_f}}{n_f!} e^{-\lambda_0} = \frac{\lambda_0^{n_m+n_f}}{n_m! n_f!} e^{-2\lambda_0}.$$

**Exercise.** Show that the **maximum likelihood estimate** for the likelihood function above is

$$\hat{\lambda}_0 = \frac{n_m + n_f}{2}.$$

Thus,

$$L(\hat{\lambda}_0 | n_m, n_f) = \frac{((n_m + n_f)/2)^{n_m+n_f}}{n_m! n_f!} e^{-(n_m+n_f)}.$$



# Chromosomal Recombination

The **likelihood ratio**, after canceling the factorial and exponential factors, is

$$\Lambda(n_m, n_f) = \frac{L(\hat{\lambda}_0 | n_m, n_f)}{L(\hat{\lambda}_m, \hat{\lambda}_f | n_m + n_f)} = \frac{(n_m + n_f)^{n_m + n_f}}{2^{n_m + n_f} n_m^{n_m} n_f^{n_f}}.$$

Our **data** for two individuals sharing the same parents are  $n_m = 56$  and  $n_f = 107$ .

Thus,

$$-2 \ln \Lambda(n_m, n_f) = -2((n_m + n_f)(\ln(n_m + n_f) - \ln 2) - n_f \ln n_f - n_m \ln n_m) = 16.228.$$

To compute the **p-value**, we first determine the number of degrees of freedom.

# Chromosomal Recombination

For this circumstance,

$$\text{degrees of freedom} = \dim(\Theta) - \dim(\Theta_0) = 2 - 1 = 1.$$

Thus we have the R commands,

```
> nm<-56
> nf<-107
> n<-nm+nf
> 1-pchisq(-2*(n*(log(n)-log(2))-nf*log(nf)-nm*log(nm)),1)
[1] 5.615274e-05
```

This very low  $p$ -value, 0.0056%, allow us to reject the null hypothesis.